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NAGAHATA On Markov processes

Suppose that  $(X_n)_{n\geq 0}$  is a Markov process. If  $(X_n)_{n\geq 0}$  satisfies

$$P(X_{n+1} = y | X_n = x) = P(X_1 = y | X_0 = x), \quad \forall n$$

then we call time homogeneous Markov process. In this lecture we only consider time homogeneous Markov process.

Let  $(X_n)_{n\geq 0}$  be a time homogeneous Markov process. We set a (huge size) square matrix

$$P_{x,y} := P(X_1 = y | X_0 = x)$$

Then we call  $P = (P_{x,y})$  transition matrix or transition probability matrix.

#### matrix formulation?

We only consider  $P(X_1 = y | X_0 = x)$ . Is it possible to formulate the conditional probability  $P(X_{n+m} = y | X_m = x)$  by means of  $P_{x,y}$ ?

We have following identities;

$$P(X_{n+m} = y | X_m = x)$$

$$= P(X_{n+m} = y, \bigcup_{y_{n+m-1}} X_{n+m-1} = y_{n+m-1} | X_m = x)$$

$$= \sum_{y_{n+m-1}} P(X_{n+m} = y, X_{n+m-1} = y_{n+m-1} | X_m = x)$$

$$= \sum_{y_{n+m-1}} P(X_{n+m} = y | X_{n+m-1} = y_{n+m-1})$$

$$\times P(X_{n+m-1} = y_{n+m-1} | X_m = x)$$

$$= \sum_{y_{n+m-1}} P(X_{n+m-1} = y_{n+m-1} | X_m = x) P_{y_{n+m-1},y}$$

Inductively, we have

$$P(X_{n+m} = y | X_m = x)$$

$$= \sum_{y_{n+m-1}} P(X_{n+m-1} = y_{n+m-1} | X_m = x) P_{y_{n+m-1},y}$$

$$\vdots$$

$$= \sum_{y_{m+1}} \sum_{y_{m+2}} \cdots \sum_{y_{n+m-1}} P_{x,y_{m+1}} P_{y_{m+1},y_{m+2}} \cdots P_{y_{n+m-1},y}$$

$$= (P^n)_{x,y}$$

Namely the conditional probability  $P(X_{n+m} = y | X_m = x)$  is given by *n*-th power of *P*.

## Definition (a random mapping representation)

A random mapping representation of transition matrix P is a function  $f: S \times \Lambda \rightarrow S$  for some  $\Lambda$  and a random variable Z with state space  $\Lambda$  and probability law Q such that it satisfies

$$Q({f(x,Z) = y}) = P_{x,y}$$

#### Proposition (a random mapping representation)

Suppose that P be a transition matrix of some time homogeneous Markov process with finite state space S. Then we have a random mapping representation f and Z.

This Proposition is one of the most important Proposition, when we write a program of some computer simulation. In particular, we write a program along the proof of this Proposition.

It is natural and applicable for us to set  $\Lambda = [0, 1]$  and Z an uniform random variable on [0, 1].

We can extend this Proposition such that we exchange a finite state space into countable state space (or much wide class). On the other hands, if we consider the application to the computer simulation, then a finite state space condition is reasonable.

<u>Proof</u> Since S is finite set, without loss of generality, we set  $S = \{s_1, s_2, \ldots, s_n\}$ . Let  $\Lambda = [0, 1]$  and Z be an uniform random variable on  $\Lambda = [0, 1]$ . We set  $\{F_{j,k}; 1 \le j \le n, 0 \le k \le n\}$  by  $F_{j,0} = 0, F_{j,k} = \sum_{i=1}^{k} P_{s_j,s_i}$ . Furthermore we set  $f : S \times \Lambda \to S$  by  $f(s_j, z) := s_k$  if  $F_{j,k-1} < z \le F_{j,k}$ . Here we have  $f(s, 0) = s_1$ . Then we can verify

$$Q(\{f(x,Z)=y\})=P_{x,y}$$

Since the matrix *P* is given by the transition probability, i.e.,  $P_{x,y} = P(X_1 = y | X_0 = x)$ , we have

$$0 \le P_{x,y} \le 1, \quad \forall x, y$$
  
 $\sum_{y} P_{x,y} = 1, \quad \forall x$ 

Conversely, if a matrix P satisfies these two identities, then there is a corresponding Markov process.

We have already seen that the conditional probability  $P(X_{n+m} = y | X_m = x)$  is given by the *n*-th power of a matrix *P*. Some computation, in particular, eigenvalue, eigenvector and diagonalization is one of the important subject. Suppose that we can diagonalize the transition matrix *P* as

$$Q^{-1}PQ = \Lambda, \quad P = Q\Lambda Q^{-1}$$
$$\Lambda = \begin{pmatrix} \lambda_1 & & \\ & \ddots & \\ & & \lambda_n \end{pmatrix}$$

If the all eigenvalues  $\lambda_i$  are real value, then it is easy to treat. One of the famous sufficient condition is that P is symmetric, i.e.,  $P_{x,y} = P_{y,x}$  for all x, y. We have an extension of this result as follows;

## Definition ( $\pi$ -symmetric)

An  $n \times n$  matrix P is symmetric with respect to  $\pi = (\pi_1, \pi_2, \dots, \pi_n) > 0$ , if and only if for all x, y it satisfies

$$\pi_{x}P_{x,y}=\pi_{y}P_{y,x}$$

Let  $\Pi$  and  $\sqrt{\Pi}$  be diagonal matrices defined via  $\pi$  by

$$\Pi = \begin{pmatrix} \pi_1 & & \\ & \ddots & \\ & & \pi_n \end{pmatrix}, \quad \sqrt{\Pi} = \begin{pmatrix} \sqrt{\pi_1} & & \\ & \ddots & \\ & & \sqrt{\pi_n} \end{pmatrix}$$

and we define a matrix S by

$$S = \sqrt{\Pi} P \sqrt{\Pi}^{-1},$$

then we can have

$$S = \sqrt{\Pi} P \sqrt{\Pi}^{-1} = \sqrt{\Pi}^{-1} (\Pi P) \sqrt{\Pi}^{-1}$$

#### Problem

Prove that if *P* is symmetric with respect to  $\pi$ , then the matrix  $\sqrt{\Pi}^{-1}(\Pi P)\sqrt{\Pi}^{-1}$  is symmetric (in usual sense).

Since S is symmetric, we can diagonalize it by using real eigenvalues by  $S = R\Lambda R^{-1}$ . Then we have

$$P = \sqrt{\Pi}^{-1} S \sqrt{\Pi} = (\sqrt{\Pi}^{-1} R) \Lambda(R^{-1} \sqrt{\Pi}) = (\sqrt{\Pi}^{-1} R) \Lambda(\sqrt{\Pi}^{-1} R)^{-1} R^{-1} R^{$$

Hence the matrix P is diagonalizable and the eigenvalues of P and S coincides.

Conversely, there is a matrix P such that the all eigenvalues are real but for any  $\pi$ , P is not symmetric with respect to  $\pi$ . Example We set

$$P = \left(\begin{array}{rrr} 0 & 1 & 0 \\ 1/2 & 0 & 1/2 \\ 1/2 & 0 & 1/2 \end{array}\right)$$

Then the eigenvalues of P is 1, 0, -1/2.

It is easy to see that if you can find a pair x, y such that  $P_{x,y} \neq 0$ and  $P_{y,x} = 0$  then for any  $\pi$ , P is not symmetric with respect to  $\pi$ . We set

$$Q = \frac{1}{12} \left( \begin{array}{rrr} 2 & 8 & 2 \\ 5 & 2 & 5 \\ 5 & 2 & 5 \end{array} \right)$$

then the eigenvalues of Q is 1, 0, -1/4. But for any  $\pi$ , Q is not symmetric with respect to  $\pi$ . (We shall prove it later.)

## equivalence class, $\leftrightarrow$ , irreducible

## Definition $(\rightarrow, \leftrightarrow)$

Let P be a transition matrix, then we define relations  $\rightarrow$ ,  $\leftrightarrow$  by ; (1)  $x \rightarrow x$ . (2) if  $\exists n, x = x_0, x_1, x_2, \dots, x_n = y$  s.t.  $P_{x_{i-1}, x_i} > 0$ ,  $1 \leq \forall i \leq n$  then  $x \rightarrow y$ (3)  $x \rightarrow y$  and  $y \rightarrow x$  then  $x \leftrightarrow y$ .

#### Problem

Show that the condition (2) is equivalent to  $\exists n \text{ s.t.}$  $P(X_n = y | X_0 = x) > 0$ .

#### Lemma (equivalence relation $\leftrightarrow$ )

The relation  $\leftrightarrow$  is an equivalence relation, namely it satisfies (1)  $x \leftrightarrow x$ (2) if  $x \leftrightarrow y$  then  $y \leftrightarrow x$ (3) if  $x \leftrightarrow y$  and  $y \leftrightarrow z$  then  $x \leftrightarrow z$ 

### Problem

It seems obvious, but prove this lemma.

#### Definition(irreducible)

If it satisfies  $\forall x, y \in S$  we have  $x \leftrightarrow y$  then we call that the state space S is irreducible or the (transition) matrix P is irreducible.

If S is not irreducible, then we can decompose S into several classes by means of the equivalence relation  $\leftrightarrow$  and we only consider following subject for each classes. Hence we only consider the irreducible case.

#### period

## Definition (period)

For each  $x \in S$  we define  $N(x) := \{n; (P^n)_{x,x} > 0\}$  and  $d_x :=$  the greatest common divisor of N(x). We call  $d_x$  period.

#### Proposition (period)

If  $x \leftrightarrow y$ , then we have  $d_x = d_y$ 

# $\frac{\text{Proof}}{\text{Since } x \leftrightarrow y, \text{ we have } x \rightarrow y \text{ and } y \rightarrow x, \text{ namely}}$

$$\exists n_1, n_2 \text{ s.t. } P(X_{n_1} = y | X_0 = x) > 0, \ P(X_{n_2} = x | X_0 = y) > 0$$

Suppose that  $n \in N(x)$  (note that we allow n = 0) then

$$P(X_n=x|X_0=x)>0.$$

Hence we have

$$P(X_{n_1+n+n_2} = y | X_0 = y)$$
  

$$\geq P(X_{n_1+n+n_2} = y, X_{n+n_2} = x, X_{n_2} = x | X_0 = y)$$
  

$$= P(X_{n_1} = x | X_0 = y) P(X_{n_2} = x | X_0 = y) P(X_{n_2} = x | X_0 = y) > 0$$

namely, we have  $n_1 + n + n_2 \in N(y)$ .

Since we allow n = 0, we also have  $n_1 + n_2 \in N(y)$ . By the definition of  $d_y$ , we have

$$\exists k \text{ s.t. } n_1 + n_2 = kd_y$$

Similarly for any  $n \in N(x)$ , we have  $n_1 + n + n_2 \in N(y)$  and

$$\exists I \text{ s.t. } n_1 + n + n_2 = Id_y$$

Hence we have

$$n=(l-k)d_y.$$

Namely we have  $d_x \ge d_y$ .

In this argument, we can exchange the role of x and y. Hence we also have  $d_y \ge d_x$ .

#### Lemma

$$\exists n_x \text{ s.t. } \forall k \geq n_x, \ kd_x \in N(x)$$

#### Problem

Prove this lemma for  $d_x = 1$ .

Suppose that  $d_x = 1$  and #S = n. Then it seems that  $n_x \le n$ . But if we set

$$P=\left(egin{array}{ccccc} 0&1/2&1/2&0\ 1&0&0&0\ 0&0&0&1\ 1&0&0&0\end{array}
ight)$$

then  $n_3 = 5 > 4 = n$ .

#### Corollary

Suppose that S is irreducible. Then the period  $d_x$  does not depend on  $x \in S$ . Hence we set this period d. We may need to rearrange the order in S and rewrite the transition matrix P, but we have following;

∃*n* s.t.

$$P^{n} = \begin{pmatrix} P_{1} & & & \\ & P_{2} & & \\ & & \ddots & \\ & & & P_{d} \end{pmatrix}$$

namely, we can decompose  $P^n$  into d block matrix. Furthermore for all  $1 \le i \le d$ ,  $P_i > 0$  (all entries of  $P_i$  is positive). Furthermore for all  $1 \le i \le d$ ,  $P_i$  are transition matrices.

Note that the matrix sizes of  $P_i$  in this corollary may not be the same, for example following is the case such that d = 2 and n = 2.

$$P = \begin{pmatrix} 0 & 0 & 1 \\ 0 & 0 & 1 \\ 1/2 & 1/2 & 0 \end{pmatrix}, P^2 = \begin{pmatrix} 1/2 & 1/2 & 0 \\ 1/2 & 1/2 & 0 \\ 0 & 0 & 1 \end{pmatrix}$$

参考文献

少なくとも以下の教科書などは参考にしている。但し、ずいぶん 前に読んで勉強した教科書の内容をそれと認識せずに参考にして いることもありえるのでその点に関してはご容赦頂きたい。

Markov 過程の基本的なところでは [K] の教科書の該当部分を多く 参考にしている。

ランダムナイトに関しては [AF] の教科書からアイデアを得ているが、証明は [K] の教科書に沿っている。

Perron-Frobenius の定理は [V] の教科書の証明に沿っているが、 確率論的には [Sai][Sal][Se] の方がよいかもしれない。なお [Sai] は 非負の行列でなく正の行列で証明をしている。

スペクトルギャップに関しては [Sal] の結果を中心に、粒子系(格子気体)への応用を述べた。粒子系への応用に関して明示的に書かれたもの(教科書、論文)はないと思う。

# 参考文献

[AF] Aldous D.; Fill J.A. Reversible Markov Chains and Random Walks on Graphs, (未刊の教科書 現在 (2015/11) も 「Aldous Fill Markov chain」で 検索をすれば pdf、 html バージョンを発見できる) [K] 小谷真一, 測度と確率, 岩波書店 [LPW] Levin D.A.; Peres Y.; Wilmer E.L. Markov Chains and Mixing Times, AMS [Sai] 斎藤正彦, 線形代数入門, 東京大学出版会 [Sal] Saloff-Coste L. Lectures on Finite Markov Chains, Lecture on Probability 1665 [Se] Seneta E. Non-negative Matrices and Markov Chains, Springer [V] Varga R.S. Matrix Iterative Analysis, Springer